ON COMONOTONICITY OF PARETO OPTIMAL RISK SHARING

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Abstract. We prove comonotonicity of Pareto-optimal risk allocations using risk measures consistent with the stochastic convex order. This extends result of Landsberger and Meilijson (1994) to risks $X \in L^1$ and general probability spaces.

1. Introduction

With the ongoing development of convex risk measures (see e.g. Föllmer and Schied (2002)) there has been renewed interest in the problem of optimal risk exchange between economic agents. A key step in studying the structure of Pareto-optimal risk allocations is the *comonotonicity* property. This result was originally obtained in Landsberger and Meilijson (1994), who provided an algorithm to construct a convex order $\leq_{cx}$-improvement of any non-comonotone allocation. Since most convex risk measures are consistent with the convex order (Bäuerle and Müller 2006) it follows that a Pareto-optimal risk allocation is necessarily comonotone. The improvement result in Landsberger and Meilijson (1994) was only stated for discrete and bounded allocations; later Dana and Meilijson (2003) constructed an extension to bounded risks on general probability spaces. In this note we further extend this idea to general risks in L^1. This is significant from a practical point of view where risks are often modeled as unbounded random variables.

2. Preliminaries

Consider the collection of real-valued integrable random variables $L^1(\mathbb{P})$ on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. As usual, we write L^∞ (resp. L^∞_+) for the collection of all bounded (positive) random variables on $(\Omega, \mathbb{P})$.

Definition 1. *Two random variables Y and $Z \in L^1(\mathbb{P})$ are said to be* comonotone *if*

$$(Y(\omega_1) - Y(\omega_2))(Z(\omega_1) - Z(\omega_2)) \geq 0, \tag{1}$$

$\mathbb{P}(d\omega_1) \times \mathbb{P}(d\omega_2)$-almost surely. In other words, Y and Z move together.

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An equivalent definition of comonotonicity is that there exist non-decreasing functions h_Y and h_Z such that $h_Y(x) + h_Z(x) = x$, $Y = h_Y(Y + Z)$, and $Z = h_Z(Y + Z)$ almost surely (Denneberg 1994).

We next recall the stochastic convex order:

Definition 2. *$Y \in L^1$ is said to* precede (*or be* preferred *to*) *$Z \in L^1$ in* convex order *if $\mathbb{E}[f(Y)] \leq \mathbb{E}[f(Z)]$ for all convex functions f for which the expectations exist. We write $Y \leq_{cx} Z$.*

Note that convex order is equivalent to ordering with respect to second stochastic dominance with equal means, see Rothschild and Stiglitz (1970, 1971).

3. Main Results

Let $X \in L^1(\mathbb{P})$ and consider the collection of integrable allocations of X, namely $\mathcal{A}(X) \triangleq \{\mathbf{Y} := (Y_1, Y_2, \ldots, Y_n) : X = \sum_{i=1}^n Y_i,\ Y_i \in L^1\}$.

Definition 3. *An allocation $\mathbf{Y} \in \mathcal{A}(X)$ is called comonotone if Y_i, Y_j are comonotone for all $1 \leq i \leq j \leq n$.*

We have trivially that if $\mathbf{Y}$ is a comonotone allocation then Y_i and X are comonotone for each i.

Theorem 1. *Let $\mathbf{Y} \in \mathcal{A}(X)$ be a discrete allocation taking on a countable number of values and bounded from below. Then there exists a comonotone allocation $\bar{\mathbf{Y}} \in \mathcal{A}(\mathbf{X})$ such that $\bar{Y}_i \leq_{cx} Y_i$, $i = 1, 2, \ldots, n$.*

The idea is to apply the Landsberger and Meilijson (1994) algorithm and take the limit using the properties of the basic single-crossing type improvement (Diamond and Stiglitz 1974). Landsberger and Meilijson (1994) proved Theorem 1 in the case where Ω was finite, i.e. when X and $\mathbf{Y}$ take on finitely many values. They further wrote that "the limiting argument [for unbounded risks X] is omitted". Indeed, the result is by now part of the folklore; however, since no complete proof has been published, and the extension is not entirely trivial, we present the entire argument below.

Proof. For completeness, we provide full details for the case $n = 2$, whence $\mathbf{Y} = (Y_1, Y_2)$. Let $\mathbf{y} = \bigcup_{i=1}^n \mathbf{y}^i$ be the union of the supports of Y_i. Then $\mathbf{y}$ is an ordered countable set in $\mathbb{R}$, bounded from below by some $\underline{\mathrm{y}}$. Consider the partition $\Omega = \cup_k C_k$ such that (Y_1, Y_2) are constant on each partition element C_k, and the C_k's are ordered according to the values of $Y_1 + Y_2$. Our goal is to have each of Y_1, Y_2 also ordered with respect to their values $y_1^1, y_2^1, \ldots, y_1^2, y_2^2, \ldots$ on C_k. Let $p_k = \mathbb{P}(C_k) > 0$ be the weights of different pairs. If Y_1 and

Y_2 are not comonotone, there must be a minimal index k such that without loss of generality $y_1^1 \le y_2^1 \le \ldots \le y_k^1 \le y_{k+1}^1$ and $y_1^2 \le y_2^2 \le \ldots \le y_k^2$, but $y_{k+1}^2 < y_k^2$. Let $1 \le j \le k$ be the minimal index such that $y_{k+1}^2 < y_j^2$, so that the $(k+1)$-st pair "violates" the comonotinicity with all pairs between j and k.

Construct the update $(\tilde{Y}_1, \tilde{Y}_2)$ which takes on the same values $\tilde{y}_i^1 \leftarrow y_i^1, \tilde{y}_i^2 \leftarrow y_i^2$ as (Y_1, Y_2) except

$$\begin{cases} \tilde{y}_{k+1}^1 & \leftarrow y_{k+1}^1 - (y_j^2 - y_{k+1}^2) \cdot (\sum_{i=j}^k p_i)/(\sum_{i=j}^{k+1} p_i), \\ \tilde{y}_{k+1}^2 & \leftarrow y_{k+1}^2 + (y_j^2 - y_{k+1}^2) \cdot (\sum_{i=j}^k p_i)/(\sum_{i=j}^{k+1} p_i), \\ \tilde{y}_i^1 & \leftarrow y_i^1 + (y_j^2 - y_{k+1}^2) \cdot p_{k+1}/(\sum_{i=j}^{k+1} p_i), \qquad j \le i \le k, \\ \tilde{y}_i^2 & \leftarrow y_i^2 - (y_j^2 - y_{k+1}^2) \cdot p_{k+1}/(\sum_{i=j}^{k+1} p_i), \qquad j \le i \le k. \end{cases} \tag{2}$$

This change preserves $\tilde{Y}_1 + \tilde{Y}_2 = Y_1 + Y_2$, maintains $\mathbb{E}\tilde{Y}_1 = \mathbb{E}Y_1$ and $\mathbb{E}\tilde{Y}_2 = \mathbb{E}Y_2$ and is an improvement of the single-crossing type (Diamond and Stiglitz 1974), also known as mean-preserving spread (mps). It follows that $\tilde{\mathbf{Y}}$ is a component-wise $\le_{cx}$ and $\le_{mps}$ improvement of $\mathbf{Y}$. Moreover, after the change, $\tilde{y}_j^2 = \tilde{y}_{k+1}^2$ and $\tilde{y}_j^1 \le \tilde{y}_{k+1}^1$ so that the j-th pair is now comonotone with the $k+1$-st pair. Therefore, the next violation index pair (k, j) will be larger (in the lexicographic order) than the current one.

Iterate this argument to obtain an improvement sequence $(\tilde{\mathbf{Y}}^{(m)})_{m=1,2,\ldots}$ of allocations. We now claim that this sequence converges almost surely and in L^1 to some limit $\bar{\mathbf{Y}}$. Indeed, let (k_m, j_m) be the violation index of the m-th update. Then $k_m \to \infty$ and therefore $p_{k_m} \to 0$. On the other hand, on a fixed subset C_k, once $k_m > k$, the value $y_k^{i,(m)}$ can change by at most

$$\begin{aligned} |y_k^{i,(m+1)} - y_k^{i,(m)}| &\le \max_{i=1,2} |y_k^{i,(m)} - y_{k_m+1}^{i,(m)}| \cdot p_{k_m+1}/(\sum_{j=k}^{k_m+1} p_j) \\ &\le \max_i (y_k^{i,(m)} - \underline{y}) \cdot p_{k_m+1}/p_k \\ &\le \big((y_k^1 + y_k^2 - 2\underline{y})/p_k\big) \cdot p_{k_m}, \end{aligned}$$

since $y_k^{1,(m)} + y_k^{2,(m)} = y_k^1 + y_k^2$ for all m.

LUDGER: ALTERNATIVELY WE CAN USE

$$\begin{aligned} |y_k^{i,(m+1)} - y_k^{i,(m)}| &\le \max_{i=1,2} |y_k^{i,(m)} - y_{k_m+1}^{i,(m)}| \cdot p_{k_m+1}/(\sum_{j=k}^{k_m+1} p_j) \\ &\le \sum_i (y_k^{i,(m)} \cdot p_{k_m+1}/p_k + \frac{1}{p_k} \sum_i y_{k_m+1}^{i,(m)} \cdot p_{k_m+1}. \end{aligned}$$

For the first term, $p_{k_m+1} \to 0$ directly, and for the second $y^{i,(m)}_{k_m+1} \cdot p_{k_m+1} \to 0$ by the integrability of $Y^{i,(m)}$ and $k_m \to \infty$. WHICH ONE IS BETTER?

Since $p_{k_m} \to 0$, for any $\epsilon > 0$ there is M large enough such that $\sum_{j>k_M} p_j < \epsilon$ and the respective tail sum is then bounded by $\sum_{n=M}^{\infty} |y^{i,(n)}_k - y^{i,(n+1)}_k| \le ((y^1_k + y^2_k - 2\underline{y})/p_k)\epsilon$. Thus, $(Y^{i,(m)})$ converges almost surely.

Moreover, $(Y^{i,(m)})$ also converges in L^1. This is obvious if Y^i is bounded since $\sup Y^{i,(m+1)} \le Y^{i,(m)}$. Otherwise note that for a fixed threshold index k', because of the mean-preserving spread the *tail mass* is non-increasing,

$$\sum_{k \ge k'} p_k \cdot y^{i,(m)}_k \le \sum_{k \ge k'} p_k \cdot y^i_k. \tag{3}$$

For a fixed level K, the tail expectation $\mathbb{E}[Y^{i,(m)} 1_{Y^{i,(m)}>K}]$ will increase only if a point (y^i_k, p_k) is moved to the right of K as a result of e.g. the third line of (2). To slide an initial point y^i_k with mass p_k to level K requires an energy of $(K - y^i_k) \cdot p_k$. Thus, at least as much energy should be subtracted from the right of K. However, total available energy beyond K is $\max_i \mathbb{E}[Y^i 1_{Y^i>K}]$. Thus, if $(K - y^i_k)p_k > \mathbb{E}[Y^i 1_{Y^i>K}]$ then the k-th point will never contribute to $\mathbb{E}[Y^{i,(m)} 1_{Y^{i,(m)}>K}]$. Let $\underline{k}_K = \min\{k : (K - y^i_k)p_k < \mathbb{E}[Y^i 1_{Y^i>K}]\}$, be the first index that can affect the above tail expectation. Combining with (3) we obtain the uniform bound

$$\mathbb{E}[Y^{i,(m)} 1_{Y^{i,(m)}>K}] \le \sum_{k \ge \underline{k}_K} p_k y^{i,(m)}_k \le \sum_{k \ge \underline{k}_K} p_k y^i_k.$$

Finally, as $K \to \infty$, $\underline{k}_K \to \infty$ and $\sum_{k \ge \underline{k}_K} p_k y^i_k = \mathbb{E}[Y^i 1_{Y^i>K}] \to 0$, establishing the uniform integrability of $(Y^{i,(m)})_m$.

The limiting $\bar{\mathbf{Y}}$ is comonotone, since there are no comonotonicity violation pairs left in the limit and by (Müller and Stoyan 2002, Theorem 1.5.9), $\bar{Y}_i \le_{cx} Y_i$, $i = 1, 2, \dots, n$.

□

When the probability space Ω is non-atomic, a more direct argument is available by extending the construction in Dana and Meilijson (2003). Namely we have,

Theorem 2. *Let $\mathbf{Y} \in \mathcal{A}(X)$ be an allocation of $X \in L^1$. Suppose Ω is non-atomic. Then there exists a comonotone allocation $\bar{\mathbf{Y}} \in \mathcal{A}(X)$ such that*

$$\bar{Y}_i \quad \le_{cx} \quad Y_i \tag{4}$$

Proof. Recall that Dana and Meilijson (2003) proved the $\le_{cx}$-improvement result for arbitrary $X \in L^\infty_+$. Defining $Y^{(m)} \triangleq \sum_{i=1}^{n} Y^{(m)}_i$, $Y^{(m)}_i = Y_i \cdot 1_{|Y_i| \le m}$, then

$$Y^{(m)} \to Y, \quad Y^{(m)}_i \to Y_i \quad \text{a.s. and in } L^1.$$

Further, by Dana and Meilijson (2003) there exists $\mathbf{Z}^{(m)}$ comonotone, such that

$$Z_i^{(m)} \leq_{cx} Y_i^{(m)} \quad \text{and} \quad \sum_{i=1}^n Z_i^{(m)} = Y^{(m)}.$$

Let $F_{i,m}$ be the distribution function of $Z_i^{(m)}$. Since $\mathbf{Z}^{(m)}$ is comonotone and Ω is non-atomic, it follows that there exists a $U \sim Unif(0,1)$ random variable such that $Z_i^{(m)} \stackrel{d}{=} F_{i,m}^{-1}(U)$. Note that for each i, $(Z_i^{(m)})_m$ are tight since by the convex ordering

$$\mathbb{E}|Z_i^{(m)}| \leq \mathbb{E}|Y_i^{(m)}| \leq \mathbb{E}|Y_i| < \infty. \tag{5}$$

Therefore there exist a subsequence, again labelled $(m) \subset \mathbb{N}$ along which the distribution functions converge, $F_{i,m} \to F_i$. This implies that $F_{i,m}^{-1}(U) \xrightarrow{a.s.} F_i^{-1}(U) =: Z_i$, and moreover

$$\sum_{i=1}^n Z_i^{(m)} \stackrel{d}{=} \sum_{i=1}^n F_{i,m}^{-1}(U) \to \sum_{i=1}^n Z_i \quad \text{a.s. and in } L^1.$$

On the other hand, we already had

$$\sum_{i=1}^n Z_i^{(m)} = \sum_{i=1}^n Y_i^{(m)} \to \sum_{i=1}^n Y_i \quad \text{a.s. and in } L^1(\mathbb{P}).$$

Thus we obtain

$$\sum_{i=1}^n Z_i \stackrel{d}{=} \sum_{i=1}^n Y_i, \quad Z_i \leq_{cx} Y_i \quad \text{and in fact} \quad (Z_i^{(m)}) \stackrel{d}{=} (F_{i,m}^{-1}(U)) \to (Z_i) \quad \text{a.s.}$$

In particular, the limit allocation $\mathbf{Z}$ is comonotone. Therefore, for some measure preserving random variable U' it holds

$$\sum_{i=1}^n Z_i \circ U' = \sum_{i=1}^n Y_i \quad \text{a.s.}$$

Obviously $\tilde{Z}_i := Z_i \circ U'$ satisfy

$$(\tilde{Z}_i) \stackrel{d}{=} (Z_i), \quad \sum_{i=1}^n \tilde{Z}_i = \sum_{i=1}^n Y_i. \quad \text{and} \quad \tilde{Z}_i \leq_{cx} Y_i. \tag{6}$$

□

Remark 1. As in Theorem 1, the $\leq_{cx}$ improvement in (5) is in fact a mean-preserving spread $\leq_{mps}$ improvement. However, the presence of $\leq_{mps}$ improvements in the above algorithm is not enough to establish tightness of $(Z_i^{(m)})$. As a counterexample, consider an allocation of a risk $X = Y_1 + Y_2$ where X and Y_1 are a mixture of a point mass at zero of weight $e > 0$ and a heavy–tailed random variable with $\mathbb{P}(X > k) = ck^{-1/4}$ for x large enough.

INSERT LUDGER'S COUNTEREXAMPLE HERE.

Therefore, the requirement of being in L^1 is necessary to reach the conclusion of Theorem 2.

4. Application to Risk Sharing

We now interpret X as the total exposure of the n agents, and use convex risk measures ρ_i to define the subjective valuation (preference) functional of the i-th agent. Thus, $\rho_i \colon L^1(\mathbb{P}) \to \mathbb{R} \cup \{+\infty\}$, are the numerical translation-invariant representations of risk-preferences of agents with $\rho_i(X) \geq \rho_i(Y)$ meaning that agent i finds X riskier than Y. We make the following assumptions on ρ_i:

Assumption 1. *ρ_i is consistent with the natural order of $L^1(\mathbb{P})$: if $\mathbb{P}(X \geq Y) = 1$ then $\rho_i(X) \geq \rho_i(Y)$.*

Assumption 2. *ρ_i is consistent with the convex order: if $X \leq_{cx} Y$ then $\rho_i(X) \leq \rho_i(Y)$.*

Assumption 3. *ρ_i is continuous with respect to a.s. convergence: if $Y^{(n)} \xrightarrow{a.s.} Y$ and $\sup_n \rho_i(Y^{(n)}) < \infty$ then $\rho_i(Y^{(n)}) \to \rho_i(Y)$. In particular, ρ_i is continuous at $-\infty$:*

$$\lim_{d \to -\infty} \rho_i[\max(Y, d)] = \rho_i(Y), \qquad \text{for all } Y \in L^1(\mathbb{P}).$$

Assumption 3 implies that given $Y \in L^1(\mathbb{P})$ and $\epsilon > 0$, there exists $\tilde{Y}$ taking on a countable number of values and bounded from below such that $|\rho_i(\tilde{Y}) - \rho_i(Y)| \leq \epsilon$. For instance, one may take $\tilde{Y} = \max(-d, 2^{-K}(\lceil 2^K Y \rceil))$ for d, K large enough.

Remark 2. One usually defines risk measures on the space $L^\infty(\mathbb{P})$, in which case an extensive theory is available, see e.g. Föllmer and Schied (2002). We also do not directly require ρ_i to be convex; see (Burgert and Rüschendorf 2006, Prop 2.2b) for relationship between convexity of ρ and its consistency with $\leq_{cx}$.

The risk sharing problem (see Aase (2002) for a recent survey) consists in finding an *optimal* allocation $\mathbf{Y}^* \in \mathcal{A}(X)$, namely an allocation such that $\mathbf{Y}^*$ is *Pareto optimal*, that is, no agent can be made strictly better off (in the sense of ρ_i-risk reduction) without another agent being made strictly worse off. Formally, a Pareto optimal risk exchange is defined as follows:

Definition 4. $\mathbf{X}^* \in \mathcal{A}(X)$ *is called a* Pareto optimal *risk exchange or allocation if whenever there exists an allocation $\mathbf{Y} \in \mathcal{A}(X)$ such that $\rho_i(Y_i) \leq \rho_i(X_i^*)$ for all $i = 1, 2, \ldots, n$, then $\rho_i(Y_i) = \rho_i(X_i^*)$ for all $i = 1, 2, \ldots, n$.*

To avoid trivialities we further restrict $\mathcal{A}(X)$ such that admissible allocations carry finite risk for each agent:

$$\mathcal{A}(X) \triangleq \{\mathbf{Y} := (Y_1, Y_2, \ldots, Y_n) : X = \sum_{i=1}^{n} Y_i,\ \rho_i(Y_i) < \infty\},$$

and assume that the new $\mathcal{A}(X)$ is still non-empty.

The improvement results in this note show that in most situations Pareto-optimal allocations are necessarily comonotone, and so one may immediately restrict the attention to comonotone allocations. The latter fact strongly simplifies the structure of the problem. For instance, in Jouini et al. (2005), Ludkovski and Young (2007) finding Pareto-optimal allocations is reduced to minimizing the linear function $f(\mathbf{Y}) := \sum_{i=1}^n \alpha_i \rho_i(Y_i)$ over possible Lagrange multipliers $\alpha_i \geq 0$. Knowing that the $\mathbf{Y}$'s of interest are comonotone allows one to work directly with the non-decreasing 1-Lipschitz functions h_i that satisfy $h_i(X) = Y_i, h_1(x) + \ldots + h_n(x) = x$.

Theorems 1 and 2 imply that once ρ_i satisfy Assumptions 1-3, then without loss of generality one can indeed perform this minimization just over the set $\mathcal{C}(X) \triangleq \{\mathbf{Y} \in \mathcal{A}(X) : \mathbf{Y} \text{ comonotone}\}$ of comonotone allocations:

$$\inf_{\mathbf{Y} \in \mathcal{A}(X)} f(\mathbf{Y}) = \inf_{\mathbf{Y} \in \mathcal{C}(X)} f(\mathbf{Y}). \tag{7}$$

When Ω is non-atomic, (7) follows immediately from Theorem 2 combined with Assumption 2. In general, one can use the improvement of Theorem 1 together with Assumption 3 to obtain an ϵ-improvement with respect the risk measures:

Lemma 1. *If $\mathbf{Y} \in \mathcal{A}(X)$, then for any $\epsilon > 0$ there is a comonotone allocation $\bar{\mathbf{Y}} \in \mathcal{C}(X)$ such that*

$$\rho_i(\bar{Y_i}) \leq \rho_i(Y_i) + \epsilon, \tag{8}$$

for all $i = 1, 2, \ldots, n$.

Clearly Lemma 1 is equivalent to (7). The proof is given in the Appendix. Note that one cannot directly "take the limit" in the discretization of Theorem 1 since $\leq_{cx}$ is not necessarily stable under a.s.-limits.

An example of a family satisfying the above assumptions are the *distortion risk measures* (alternatively known as the law-invariant, comonotone-additive coherent risk measures). Denote by S_Y the (decumulative) distribution function of Y, that is, $S_Y(t) = \mathbb{P}(Y > t)$, and by S_Y^{-1} the (pseudo-)inverse of S_Y, which is unique up to a countable set (Denneberg 1994). For concreteness, take $S_Y^{-1}(p) = \sup\{t : S_Y(t) > p\}$; the inverse S_Y^{-1} thus defined is right continuous. Let $g : [0,1] \to [0,1]$ be a non-decreasing, concave function such that $g(0) = 0$, $g(1) = 1$. Take

$$\begin{aligned} \rho(Y) = \rho_g(Y) = \int Y \, d(g \circ \mathbb{P}) = \int_0^1 S_Y^{-1}(p) \, dg(p) \qquad (9) \\ = \int_{-\infty}^0 \left(g[S_Y(t)] - 1\right) \, dt + \int_0^\infty g[S_Y(t)] \, dt. \end{aligned}$$

The function g is called a distortion because it modifies, or distorts, the tail probability S_Y before calculating an expectation. Note that if $g(p) = p$, then $\rho_g(Y) = \mathbb{E}Y$. For this reason, ρ_g is sometimes referred to an expectation with respect to a distorted probability.

Recall that for concave g, ρ_g is monotone, continuous at $-\infty$ and consistent with $\leq_{cx}$ (Wang and Young 1998). Moreover, ρ_g is continuous with respect to a.s.-convergence. For $0 < p < 1$, let $\delta(p, \epsilon) > 0$ be such that $g(p + \eta) < g(p) + \epsilon$ for all $0 < \eta < \delta(p, \epsilon)$ (since g is uniformly continuous away from zero). Now define $\tilde{Y} \geq Y$ such that $\tilde{Y}$ takes on countably many values $\mathbf{y} = \{y_k\}_{k=1}^{\infty}$, $y_1 > -\infty$ and

$$S_Y(t) \leq S_{\tilde{Y}}(t) \leq S_Y(t) + \delta(S_Y(t), \epsilon \wedge \epsilon t^{-2}).$$

Without loss of generality, assume also that $S_{\tilde{Y}}(y_k) = S_Y(y_k)$ for all k. Because the distortion function g is concave, it is uniformly Lipschitz on $(\eta, 1]$ for any $\eta > 0$. Therefore, $\mathbf{y}$ forms a discrete set (without cluster points with respect to the usual metric on $\mathbb{R}$), i.e. $|y_{k'} - y_k| > \epsilon_2 > 0$ for any $k' \neq k$. Finally,

$$\begin{aligned}
\rho_g(Y) \leq \rho_g(\tilde{Y}) &= \int_{-\infty}^{0} [g(S_{\tilde{Y}}(t)) - 1] \, dt + \int_{0}^{\infty} g(S_{\tilde{Y}}(t)) \, dt \\
&\leq \int_{-\infty}^{0} \left[g(S_Y(t)) + (\epsilon \wedge \epsilon t^{-2}) - 1\right] dt + \int_{0}^{\infty} \left[g(S_Y(t)) + (\epsilon \wedge \epsilon t^{-2})\right] dt \\
&\leq \rho_g(Y) + 4\epsilon.
\end{aligned}$$

where the second line follows by definition of $\delta(\cdot, \cdot)$.

Appendix

Proof of Lemma 1. Consider a given allocation $\mathbf{Y}$. Fix $\epsilon > 0$. Using Assumption 3, construct a discrete $\tilde{Y}_i \geq Y_i$ such that

$$\rho_i(Y_i) \leq \rho_i(\tilde{Y}_i) \leq \rho_i(Y_i) + \epsilon.$$

Using Theorem 1, there is an $\leq_{cx}$-imporevement $\mathbf{Y}^\epsilon$ of $\tilde{\mathbf{Y}}$. Unfortunately, $\mathbf{Y}^\epsilon$ is not directly comparable with $\mathbf{Y}$ since the sum of the risks is now too big, i.e. greater than X. The remaining steps show that one can further minorize $\mathbf{Y}^\epsilon$ such that the allocation remains comonotone and the sum is simply X.

First, we construct a comonotone improvement $\mathbf{Y}^\mathbf{c}$ of $\mathbf{Y}^\epsilon$ such that $\sum_i Y_i^c$ is comonotone with X. Let $\mathbf{y}$ be the support of $\sum_i Y_i^\epsilon$. Order $\mathbf{y} = \{\ldots < y_k < y_{k+1} < \ldots\}$. Since $\mathbf{Y}^\epsilon$ is comonotone, each Y_i^ϵ must be constant on an event $\{\omega \colon \sum_i Y_i^\epsilon = y_k\}$, say $Y_i^\epsilon = b_k^i$.

Define $B_k \triangleq \{\omega \colon y_{k-1} < X(\omega) \leq y_k\}$. Clearly, $\cup_k B_k = \Omega$ is a measurable partition. Now set $\mathbf{Y}_\mathbf{i}^\mathbf{c}(\omega) = b_k^i$, whenever $\omega \in B_k$. Then $\mathbf{Y}^\mathbf{c}$ is still comonotone and $\sum_i Y_i^c \geq X$ by construction. Moreover, $\mathbf{Y}^\mathbf{c}$ and X are comonotone: consider $\omega_1 \in B_{k_1}$ and suppose

$X(\omega_1) > X(\omega_2)$. If $\omega_2 \in B_{k_2}$ then necessarily $k_2 \leq k_1$ and $\sum_i Y_i^c(\omega_1) = y_{k_1} \geq y_{k_2} = \sum_i Y_i^c(\omega_2)$. Finally, $\mathbf{Y^c} \leq \mathbf{Y}^\epsilon$, since by construction $\mathbf{Y^c}$ is the smallest random variable with support in $\mathbf{y}$ that dominates X. For later use we also assume without loss of generality that $y_k = \operatorname{ess\,sup}_{\omega \in B_k} X(\omega)$.

The final step improves the allocation to $\bar{\mathbf{Y}}$, such that $\sum_i \bar{Y}_i = X$. Let $X' = \sum_i Y_i^c$. Since X and X' are comonotone, there exists a random variable Z such that $X = h(Z), X' = h'(Z)$ for some non-decreasing functions h, h'. Since a collection of comonotone random variables is also comonotone with their sum, there exist continuous increasing functions $\tilde{f}_i$ such that $\sum_i \tilde{f}_i(x) = x$ and $Y_i^c = \tilde{f}_i(X')$. Set $f_i = \tilde{f}_i \circ h'$, so that $Y_i^c = f_i(Z)$. Recall that each Y_i^c is discrete, so that f_i is piecewise constant and non-decreasing. The last step of the proof shows that one can construct non-decreasing $\bar{f}_i$'s such that $\bar{f}_i \leq f_i$ and $\sum_i \bar{f}_i = h \leq h' = \sum_i f_i$, which is geometrically intuitive.

Partition $\mathbb{R}$ into disjoint intervals $C_k \triangleq (z_k, z_{k+1})$ such that $\sum_i f_i$ is constant on each interval. Then without loss of generality there is an increasing sequence (a_k) such that on each interval, $a_k \leq h(z) \leq a_{k+1}$, $z_k \leq z \leq z_{k+1}$. Take $r_k = \sup\{z \in C_k \colon h(z) \leq (a_k + a_{k+1})/2\}$ to be the lower "half" of each C_k. By construction, $\sum_i b_k^i = \sup_{z \in C_k} h(z) = a_{k+1}$. Set $d_k^i = (b_{k+1}^i - b_k^i) \geq 0$, so that $\sum_i d_i^k = a_{k+1} - a_k > 0$. Define $\bar{f}_i(z) = b_k^i - \eta_k^i$ if $a_k \leq z \leq r_k$ and $\bar{f}_i(z) = b_k^i$ if $r_k < z \leq a_{k+1}$. Observe that the new family $\bar{f}_i$ is still non-decreasing, $\sum_i f_i \geq \sum_i \bar{f}_i(z) \geq h$ and $\| \sum_i \bar{f}_i - h\|_\infty \leq \frac{1}{2} \| \sum_i f_i - h\|_\infty$. This construction is illustrated in Figure 1. By repeating the argument, we obtain a monotonically decreasing sequence of comonotone allocations that converges pointwise. With a slight abuse of notation, let $\bar{\mathbf{Y}} = \bar{f}(Z)$ be the limiting allocation. Then the sum of $\bar{f}_i$ is h (i.e. $\sum_i \bar{Y}_i = X$), and $\bar{\mathbf{Y}}$ is still comonotone since each $\bar{f}_i$ is non-decreasing.

To conclude the proof observe that $\bar{\mathbf{Y}}$ is a comonotone allocation of X, and

$$\rho_i(\bar{Y}_i) \leq \rho_i(Y_i^\epsilon) \leq \rho_i(\tilde{Y}_i) \leq \rho_i(Y_i) + 4\epsilon, \tag{10}$$

for all $i = 1, 2, \ldots, n$, matching (8). □

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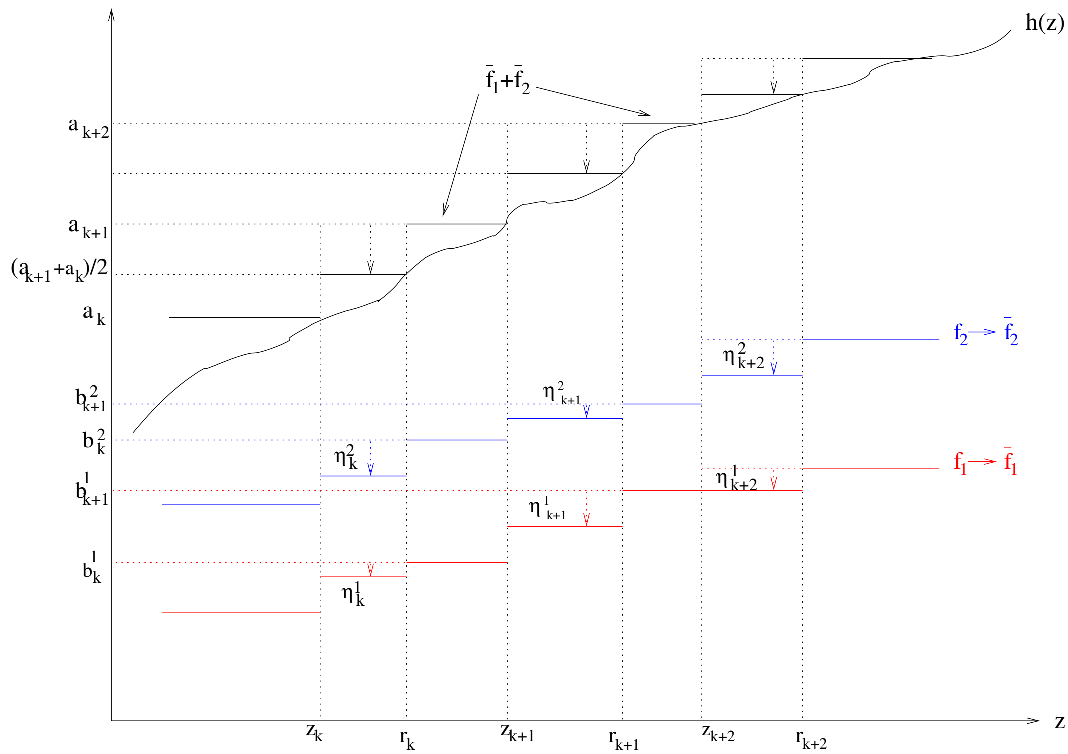

FIGURE 1. Third Step of Lemma 1 that improves the comonotone allocation described by (f_i) to $(\bar{f}_i)$ for the case $n = 2$.

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