

**CONSERVATION OF THE UMP-RESP. MAXIMIN-PROPERTY  
OF STATISTICAL TESTS UNDER EXTENSIONS OF  
PROBABILITY MEASURES\***

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**1. INTRODUCTION**

Let  $(X, \mathcal{A})$  be a measure space, let  $\mathcal{A}_0 \subset \mathcal{A}$  be a sub  $\sigma$ -algebra and let  $\mathcal{P}_0, \mathcal{P}_1$  be two sets of probability measures on  $\mathcal{A}_0$ . In the present paper we consider testproblems  $\mathcal{P}'_0, \mathcal{P}'_1$ , where  $\mathcal{P}'_i$  are subsets of the extensions of elements of  $\mathcal{P}_i$  to the larger  $\sigma$ -algebra  $\mathcal{A}$ . Especially we are interested in the question, in which cases optimality properties of tests for  $\mathcal{P}_0, \mathcal{P}_1$  can be lifted to the testproblem  $\mathcal{P}'_0, \mathcal{P}'_1$ . (We consider maximin-tests and UMP-tests only.) Although this question looks a little bit artificial, testing problems of this kind occur in many practical situations. Some examples are the following:

(a) Let  $\mathcal{P}'_0, \mathcal{P}'_1$  be sets of probability measures on  $(X, \mathcal{A})$  and suppose that for some reason one only can observe a function  $T$  for testing  $\mathcal{P}'_0, \mathcal{P}'_1$ . Then defining  $\mathcal{P}_i$  to be the images of  $\mathcal{P}'_i$  under  $T$  we answer

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the question whether optimal tests for  $\mathcal{P}_0, \mathcal{P}_1$  (based on  $T$ ) are optimal for  $\mathcal{P}'_0, \mathcal{P}'_1$ .

(b) Let  $G$  be a set of transformations on  $(\mathcal{X}, \mathcal{A})$ , let  $\mathcal{A}_0$  be the  $\sigma$ -algebra of  $G$ -invariant sets, let  $\mathcal{P}_0, \mathcal{P}_1$  be probability measures on  $(\mathcal{X}, \mathcal{A})$  and consider the testproblem

$$\mathcal{P}'_0 = \{P_0^g \mid g \in G\}, \quad \mathcal{P}'_1 = \{P_1^g \mid g \in G\}.$$

Thus especially the Hunt–Stein situation is covered by our model.

(c) Let  $\mathcal{P}'_i = \{P_{(\theta, \eta)}; \theta \in \Theta_i, \eta \in \Gamma\}$ ,  $i = 0, 1$ , be two sets of probability measures on  $(\mathcal{X}, \mathcal{A})$ . Consider the test problem  $\Theta_0 : \Theta_1$ ; thus in this case  $\eta \in \Gamma$  is a nuisance parameter. Assume that  $\mathcal{A}_0 \subset \mathcal{A}$  is a sub  $\sigma$ -algebra such that the restriction of  $P_{(\theta, \eta)}$  on  $\mathcal{A}_0$  is independent of  $\eta \in \Gamma$ . So the question is whether  $\mathcal{A}_0$  is "sufficient" for the test problem  $\Theta_0 : \Theta_1$  in the presence of nuisance parameters.

In Section 2 of this paper we derive some properties of extensions of probability measures. In Section 3 we consider the question whether the maximin-property of tests for  $\mathcal{P}_0, \mathcal{P}_1$  may be lifted. We also discuss this question in connection with the concept of least favourable distributions. In Section 4 we consider UMP-tests and finally in Section 5 we give some examples.

## 2. SOME PROPERTIES OF EXTENSIONS OF PROBABILITY MEASURES

Let  $(\mathcal{X}, \mathcal{A})$  be a measure space,  $M_1(\mathcal{X}, \mathcal{A})$  be the set of probability measures on  $(\mathcal{X}, \mathcal{A})$  and  $\mathcal{A}_0 \subset \mathcal{A}$  be a sub  $\sigma$ -algebra of  $\mathcal{A}$ . For  $P \in M_1(\mathcal{X}, \mathcal{A}_0)$  let  $E(P) \subset M_1(\mathcal{X}, \mathcal{A})$  denote the set of all extensions of  $P$  to the larger  $\sigma$ -algebra  $\mathcal{A}$ . Furthermore, for  $\mathcal{P} \subset M_1(\mathcal{X}, \mathcal{A}_0)$  let  $E(\mathcal{P}) = \bigcup_{P \in \mathcal{P}} E(P)$ .

The aim of this section is to present some properties of the set  $E(\mathcal{P})$  of extensions. In the following we shall always consider the relative weak  $*$ -topology on  $M_1(\mathcal{X}, \mathcal{A})$ . The first proposition is trivial.

**Proposition 1.**

- (a) If  $\mathcal{P} \subset M_1(X, \mathcal{A}_0)$  is convex, then  $E(\mathcal{P})$  is convex.  
 (b)  $E(\overline{\mathcal{P}}) \supset \overline{E(\mathcal{P})}$  ( $\overline{A}$  denoting the closure of  $A$ ).

Generally  $E(\overline{\mathcal{P}}) \neq \overline{E(\mathcal{P})}$  (consider for example  $X = [0, \Omega)$ ,  $\Omega$  first uncountable ordinal number,  $\mathcal{A}_0$   $\sigma$ -algebra generated by  $[0, \alpha)$ ,  $0 \leq \alpha < \Omega$ ,  $\mathcal{A}$  generated by  $\mathcal{A}_0 \cup \{\Omega\}$ , and  $\mathcal{P} = \{\delta_\alpha \mid 0 \leq \alpha < \Omega\}$ ), and for  $P \in M_1(X, \mathcal{A}_0)$  the set  $E(P)$  may be empty. (For a discussion of this point cf. Lipiecki [8].) A simple condition for  $E(P) \neq \phi$  and an interesting extension is given by the following proposition. Let  $M(X, \mathcal{A})$  denote the set of all measures  $\mu$  on  $(X, \mathcal{A})$  such that  $\mu|_{\mathcal{A}_0}$  is  $\sigma$ -finite.

**Proposition 2.** For  $\mu \in M(X, \mathcal{A})$  and  $P \ll \mu|_{\mathcal{A}_0}$  define  $P_\mu(A) = \int_A \frac{dP}{d\mu|_{\mathcal{A}_0}} d\mu$ ,  $A \in \mathcal{A}$ , and  $(E(P))_\mu = \{Q \in E(P) \mid Q \ll \mu\}$ . Then it holds:

- (a)  $P_\mu \in E(P)$ ,  
 (b)  $Q \ll P_\mu$  for  $Q \in (E(P))_\mu$ .

**Proof.**

(a) is immediate by definition.

(b) If  $A \in \mathcal{A}$ ,  $P_\mu(A) = 0$ , then  $1_A \frac{dP}{d\mu|_{\mathcal{A}_0}} = 0[\mu]$  and, therefore,  $1_A \frac{dP}{d\mu|_{\mathcal{A}_0}} = 0[Q]$ . Furthermore,  $\frac{dP}{d\mu|_{\mathcal{A}_0}} > 0[P]$  implies that  $\frac{dP}{d\mu|_{\mathcal{A}_0}} > 0[Q]$  and, therefore,  $1_A = 0[Q]$ .

$P_\mu$  is the uniquely determined  $\mu$ -continuous extension which has a  $\mathcal{A}_0$ -measurable density. The set of all  $\mu$ -continuous extensions is given by the following proposition. Let  $L(\mathcal{A})$  denote the  $\mathcal{A}$ -measurable real functions.

**Proposition 3.**  $\mu \in M(X, \mathcal{A})$  and  $P \ll \mu|_{\mathcal{A}_0}$  implies  $E(P)_\mu = \{h\mu \mid h \in L(\mathcal{A}), h \geq 0, E_\mu(h|_{\mathcal{A}_0}) = \frac{dP}{d\mu|_{\mathcal{A}_0}}\}$ .

**Proof.** If  $Q \in E(P)_\mu$ , then by the theorem of Radon–Nikodym  $Q = h\mu$ ,  $h \in L(\mathcal{A})$ ,  $h \geq 0$ , and for  $A_0 \in \mathcal{A}_0$ ,  $Q(A_0) = \int_{A_0} h \, d\mu = P(A_0) = \int_{A_0} \frac{dP}{d\mu | \mathcal{A}_0} \, d\mu$ . Therefore,  $E_\mu(h | \mathcal{A}_0) = \frac{dP}{d\mu | \mathcal{A}_0} [\mu | \mathcal{A}_0]$ . If  $E_\mu(h | \mathcal{A}_0) = \frac{dP}{d\mu | \mathcal{A}_0}$ , then  $\int_{A_0} h \, d\mu = \int_{A_0} E_\mu(h | \mathcal{A}_0) \, d\mu = P_\mu(A_0) = P(A_0)$  and, therefore,  $h\mu \in E(P)_\mu$ .

**Corollary 1.** Let  $Q \in E(P)$ ,  $f \in L_1(\mathcal{A}, Q)$  with  $f \geq 0$  and assume  $E_Q(f | \mathcal{A}_0) > 0$ . Define  $h = \frac{f}{E_Q(f | \mathcal{A}_0)}$  and  $Q^{(\wedge)} = hQ$ , then  $Q^{(\wedge)} \in E(P)$ .

A consequence of Corollary 1 is, that if there is any extension of  $P$ , then there are many extensions. The following proposition expresses this fact in statistical terms: If a test  $\varphi \in \Phi$  is similar w.r.t.  $E(P)_Q$  then  $\varphi$  is  $Q$ -a.s.  $\mathcal{A}_0$ -measurable, where  $\Phi$  denotes the set of all tests.

**Proposition 4.** Let  $Q \in E(P)$ ,  $\varphi \in \Phi$  and  $E_Q\varphi = E_{Q^{(\wedge)}}\varphi$  for all  $f \in L_1(\mathcal{A}, Q)$  with  $f \geq 0$  and  $E_Q(f | \mathcal{A}_0) > 0$ . Then  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$ .

**Proof.** Because of  $A_0 := \{E_Q(\varphi | \mathcal{A}_0) = 0\} \subset \{\varphi = 0\}[Q]$  one obtains  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$  on  $A_0$  and

$$(*) \quad \int \varphi \, dQ = \int_{A_0^c} E_Q(\varphi | \mathcal{A}_0) \, dQ = \int_{A_0^c} \frac{E_Q(\varphi f | \mathcal{A}_0)}{E_Q(f | \mathcal{A}_0)} \, dQ.$$

Choosing  $f = \varphi + \delta$ ,  $\delta > 0$  one gets by the lemma of Fatou for  $\delta \rightarrow 0$

$$\int_{A_0^c} E_Q(\varphi | \mathcal{A}_0) \, dQ \geq \int_{A_0^c} \frac{E_Q(\varphi^2 | \mathcal{A}_0)}{E_Q(\varphi | \mathcal{A}_0)} \, dQ.$$

By Jensen's inequality  $E_Q(\varphi^2 | \mathcal{A}_0) \geq E_Q^2(\varphi | \mathcal{A}_0)[Q]$ , which yields  $E_Q(\varphi^2 | \mathcal{A}_0) = E_Q^2(\varphi | \mathcal{A}_0)[Q]$  on  $A_0^c$  implying  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$  on  $A_0^c$ .

Together, we have  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$ .

The extreme points of  $E(P)$  are characterized by a theorem due to Douglas [3]. Our special extensions of Proposition 2 allow to prove

(a slight modification of) Douglas' result by probabilistic methods without using the Hahn-Banach theorem.

**Proposition 5.** Let  $Q \in E(P)$ . Then  $Q$  is an extreme point of  $E(P)$  if and only if it holds that  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$  for all  $\varphi \in \Phi$  or, equivalently, if for all  $A \in \mathcal{A}$  there is a  $B \in \mathcal{A}_0$  with  $Q(A \Delta B) = 0$ .

**Proof.** If  $Q = \alpha Q_1 + (1 - \alpha)Q_2$ ,  $Q_i \in E(P)$ ,  $i = 1, 2$ , then  $Q_1 \ll Q$  and  $\frac{dQ_1}{dQ} \leq \frac{1}{\alpha}$ ,  $\alpha \in (0, 1)$ . From  $\frac{dQ_1}{dQ} = E_Q\left(\frac{dQ_1}{dQ} | \mathcal{A}_0\right)[Q]$  and  $Q | \mathcal{A}_0 = Q_1 | \mathcal{A}_0$  we obtain  $Q = Q_1$ .

For  $\varphi \in \Phi$  define  $h = \frac{1 + \varphi}{E_Q(1 + \varphi | \mathcal{A}_0)} - 1$  then  $-1 \leq h \leq 1$  and by Corollary 1  $(1 + h)Q, (1 - h)Q \in E(P)$  and  $Q = \frac{1}{2}[(1 + h)Q + (1 - h)Q]$ . Therefore,  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$ .

**Remark 1.** In a similar way one could also give a simple proof of the following characterization of the extreme points of the set  $E^G(P)$  of all extensions of  $P$  which are invariant w.r.t. a semigroup  $G$ , namely: Let  $Q \in E^G(P)$ ; then  $Q$  is an extreme point of  $E^G(P)$  iff for all  $\varphi \in \Phi$  which are  $Q$ -almost  $G$ -invariant  $\varphi = E_Q(\varphi | \mathcal{A}_0)[Q]$  holds. This result is due to Luschgy [9], Theorem 4.4.

An important property of extensions gives the following proposition. For a set  $A$  in a vector space let  $\text{con}(A)$  be the convex hull of  $A$ .

**Proposition 6.**

(a) If  $P_i \in M_1(X, \mathcal{A}_0)$ ,  $E(P_i) \neq \emptyset$ ,  $i = 0, 1$ , then for all  $\alpha \in [0, 1]$

$$E(\alpha P_0 + (1 - \alpha)P_1) = \alpha E(P_0) + (1 - \alpha)E(P_1).$$

(b) If  $\mathcal{P} \subset M_1(X, \mathcal{A}_0)$  with  $E(P) \neq \emptyset$  for all  $P \in \mathcal{P}$ , then  $E(\text{con } \mathcal{P}) = \text{con } E(\mathcal{P})$ .

**Proof.**

(a) If  $Q_i \in E(P_i)$ ,  $i = 0, 1$ , then trivially,  $\alpha Q_0 + (1 - \alpha)Q_1 \in E(\alpha P_0 + (1 - \alpha)P_1)$ . Let  $Q \in E(\alpha P_0 + (1 - \alpha)P_1)$  and  $\alpha \in (0, 1)$  then  $P_0 \ll Q | \mathcal{A}_0$  and so we obtain by Proposition 2  $\tilde{P}_0(A) = \int_A \frac{dP_0}{dQ | \mathcal{A}_0} dQ$

defines an extension of  $P_0$  with  $\tilde{P}_0(A) \leq \frac{1}{\alpha} Q(A)$ . Define  $\tilde{P}_1(A) = \frac{1}{1-\alpha} (Q(A) - \alpha \tilde{P}_0(A))$ . Then  $\tilde{P}_1 \in E(P_1)$  and  $Q = \alpha \tilde{P}_0 + (1-\alpha) \tilde{P}_1$ .

(b) is immediate from (a).

**Proposition 7.** Let  $P_i \in M_1(X, \mathcal{A}_0)$ ,  $i = 0, 1$ , and  $P_0 = hP_1$  where  $h \in L_1(\mathcal{A}_0, P_1)$  with  $h \geq 0$ . If  $E(P_1) \neq \emptyset$ , then

$$E(P_0) = hE(P_1) = \{hQ \mid Q \in E(P_1)\}.$$

**Proof.** If  $Q_1 \in E(P_1)$ , then for  $A_0 \in \mathcal{A}_0$

$$(hQ_1)(A_0) = \int_{A_0} h dQ_1 = \int_{A_0} h dP_1 = P_0(A_0).$$

If  $h > 0$  [ $P_1$ ], then  $P_1 = \frac{1}{h} P_0$  and, therefore, by the first inclusion  $E(P_1) \supset \frac{1}{h} E(P_0) \supset E(P_1)$  which implies  $E(P_0) = hE(P_1)$ . Assume now  $0 < a = P_1(\{h = 0\}) < 1$  and define  $X' = X \setminus \{h = 0\}$ ,  $\mathcal{A}' = \mathcal{A} \cap X'$ ,  $\mathcal{A}'_0 = \mathcal{A}_0 \cap X'$ ,  $P'_0 = P_0 \mid \mathcal{A}'_0$ ,  $P'_1 = \frac{1}{1-a} P_1 \mid \mathcal{A}'_0$  and  $h' = h \mid X'$ . Then  $P'_0 = (1-a)h'P'_1$  and  $Q_0 \in E(P_0)$  implies  $Q'_0 = Q_0 \mid \mathcal{A}'_0 \in E(P'_0)$ . Therefore, there exists  $Q'_1 \in E(P'_1)$  such that  $Q'_0 = (1-a)h'Q'_1$ . Let  $\tilde{Q} \in E(P_1)$  and define  $Q_1(A) = (1-a)Q'_1(A \cap \{h > 0\}) + \tilde{Q}(A \cap \{h = 0\})$ . We have for  $A_0 \in \mathcal{A}_0$ :

$$Q_1(A_0) = (1-a)P'_1(A_0 \cap \{h > 0\}) + P_1(A_0 \cap \{h = 0\}) = P_1(A_0)$$

and

$$Q_0 = hQ_1.$$

Proposition 7 allows to describe the relation between extensions of two probability measures completely.

**Proposition 8.** Let  $P_0, P_1 \in M_1(X, \mathcal{A}_0)$ , let  $P_0 = aP'_0 + (1-a)P''_0$  where  $P'_0, P''_0 \in M_1(X, \mathcal{A}_0)$  and  $P'_0 = hP_1$ ,  $P''_0 \perp P_1$ ,  $a \in [0, 1]$ .

Assume that  $E(P_i) \neq \emptyset$ ,  $i = 0, 1$ . Then

$$(a) E(P'_0), E(P''_0) \neq \phi.$$

$$(b) E(P_0) = ahE(P_1) + (1-a)E(P''_0).$$

$$(c) E(P''_0) \perp E(P_1).$$

**Proof.**

(a) Since  $P'_0, P''_0 \ll P_0$  (a) is immediate from Proposition 2.

(b) By Proposition 6 and 7  $E(P_0) = aE(P'_0) + (1-a)E(P''_0) = ahE(P_1) + (1-a)E(P''_0)$ .

(c) There exists a  $A \in \mathcal{A}_0$  with  $P''_0(A^c) = 0, P_1(A) = 0$ . Therefore,  $Q(A^c) = 0$  for  $Q \in E(P''_0)$  and  $R(A) = 0$  for  $R \in E(P_1)$ . This implies that  $E(P''_0) \perp E(P_1)$ .

A wellknown criterion for sufficiency implies the following

**Corollary 2.** Let  $P_i, i = 0, 1$ , be as in Proposition 8. Let  $Q_i \in E(P_i), i = 0, 1$ . Then  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$  iff

$$Q_0 \in ahQ_1 + (1-a)E(P''_0)$$

with  $h \in L_1(\mathcal{A}_0, P_1)$  and  $h \geq 0$ .

For  $k \geq 0$  and  $P, Q \in M_1(\mathcal{X}, \mathcal{A})$  define the distances

$$d_k(Q, P) = \|Q - kP\| =$$

$$\sup \{Q(A) - kP(A) - (Q(B) - kP(B)), A, B \in \mathcal{A}\},$$

and for  $P, Q \subset M_1(\mathcal{X}, \mathcal{A})$

$$d_k(\mathcal{P}, \mathcal{Q}) = \inf \{d_k(P, Q) \mid P \in \mathcal{P}, Q \in \mathcal{Q}\}.$$

**Proposition 9.** Let  $\mathcal{P}_i \subset M_1(\mathcal{X}, \mathcal{A}_0), i = 0, 1$ , let  $E(P) \neq \phi$  for all  $P \in \mathcal{P}_0 \cup \mathcal{P}_1$  and let  $k \geq 0$ .

(a) If  $\mu \in M(\mathcal{X}, \mathcal{A}), P_0, P_1 \in M_1(\mathcal{X}, \mathcal{A}_0)$ , with  $P_i \ll \mu \mid \mathcal{A}_0, i = 0, 1$ , then

$$d_k(P_0, P_1) = d_k(P_{0,\mu}, P_{1,\mu}).$$

$$(b) d_k(\mathcal{P}_0, \mathcal{P}_1) = d_k(E(\mathcal{P}_0), E(\mathcal{P}_1)).$$

**Proof.**

(a) It is easy to see that

$$d_k(P_0, P_1) = \|P_0 - kP_1\| = \max\{k - 1 + 2(P_0 - kP_1)_+(x)\},$$

where  $(P_0 - kP_1)_+$  is the positive part of the Jordan-Hahn decomposition of  $P_0 - kP_1$ . Therefore, with

$$A = \left\{ \frac{dP_0}{d\mu|_{\mathcal{A}_0}} \geq k \frac{dP_1}{d(\mu|_{\mathcal{A}_0})} \right\}$$

$$(P_0 - kP_1)_+(x) = (P_0 - kP_1)(A) = (P_{0,\mu} - kP_{1,\mu})(A).$$

(b) For  $P_i \in \mathcal{P}_i$  and  $Q_i \in E(\mathcal{P}_i)$ ,  $i = 0, 1$ ,

$$\begin{aligned} & \sup\{|P_0(A) - kP_1(A)|; A \in \mathcal{A}_0\} = \\ & \sup\{|Q_0(A) - kQ_1(A)|; A \in \mathcal{A}_0\} \leq \\ & \sup\{|Q_0(A) - kQ_1(A)|; A \in \mathcal{A}\}. \end{aligned}$$

Thus,  $\|P_0 - kP_1\| \leq \|Q_0 - kQ_1\|$ .

By part (a) equality holds for  $Q_i = P_{i,\mu}$ ,  $i = 0, 1$ . This implies  $d_k(\mathcal{P}_0, \mathcal{P}_1) = d_k(E(\mathcal{P}_0), E(\mathcal{P}_1))$ .

### 3. MAXIMIN-TESTS AND LEAST FAVOURABLE DISTRIBUTIONS

Let  $\mathcal{P}_0, \mathcal{P}_1 \in M_1(\mathcal{X}, \mathcal{A}_0)$  and let  $\mathcal{P}'_i \subset E(\mathcal{P}_i)$ ,  $i = 0, 1$ , such that  $\overline{\text{con}} \mathcal{P}'_i \cap E(P) \neq \emptyset$ , for all  $P \in \mathcal{P}_i$ ,  $i = 0, 1$ . For the testproblem  $\mathcal{P}'_0, \mathcal{P}'_1$  denote the maximin-risk for  $\alpha \in [0, 1]$  by

$$\beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) = \sup_{\varphi \in \Phi_\alpha(\mathcal{P}'_0, \mathcal{A})} \inf_{Q \in \mathcal{P}'_1} E_Q \varphi,$$

where  $\Phi_\alpha(\mathcal{P}'_0, \mathcal{A}) = \{\varphi: (\mathcal{X}, \mathcal{A}) \rightarrow ([0, 1], [0, 1]^{\mathcal{B}^1}) \mid E_Q \varphi \leq \alpha \text{ for all } Q \in \mathcal{P}'_0\}$  and  $\mathcal{B}^1$  is the Borel  $\sigma$ -algebra.

The general assumption  $\overline{\text{con}} \mathcal{P}'_i \cap E(P) \neq \emptyset$  for  $P \in \mathcal{P}_i$ ,  $i = 0, 1$ , implies that  $\beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) \geq \beta(\alpha, \mathcal{P}_0, \mathcal{P}_1)$ . The following theorem gives a



sufficient conditions to imply that a maximin solution for  $\mathcal{P}_0, \mathcal{P}_1$  is even a solution for  $\mathcal{P}'_0, \mathcal{P}'_1$  i.e. optimality of a test for  $\mathcal{P}_0, \mathcal{P}_1$  is inherited to the testproblem  $\mathcal{P}'_0, \mathcal{P}'_1$ .

**Theorem 10.** *If  $\mathcal{P}_i \ll \mu$ ,  $i = 0, 1$ , and if  $d_k(\text{con } \mathcal{P}'_1, \text{con } \mathcal{P}'_0) = d_k(\text{con } \mathcal{P}_1, \text{con } \mathcal{P}_0)$  for all  $k \geq 0$  then  $\beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) = \beta(\alpha, \mathcal{P}_0, \mathcal{P}_1)$  for all  $\alpha \in (0, 1]$ .*

**Proof.** By Baumann [1], Satz 6.3, for a dominated testproblem  $H, K$  on  $\mathcal{X}$  one has  $\beta(\alpha, H, K) = \min\{\alpha k + (Q - kP)_+(X) \mid k \geq 0, Q \in \overline{\text{con } K}, P \in \overline{\text{con } H}\}$  (the closure w.r.t. relative weak \*-topology). Using  $(Q - kP)_+(X) = \frac{1-k}{2} + \frac{1}{2} d_k(Q, P)$  and the fact that for  $\epsilon > 0$  there is a measure  $\tilde{\mu}$  on  $(X, \mathcal{A})$  such that  $d_k((\text{con } \mathcal{P}'_0)_{\tilde{\mu}}, (\text{con } \mathcal{P}'_1)_{\tilde{\mu}}) \leq d_k(\text{con } \mathcal{P}'_0, \text{con } \mathcal{P}'_1) + \epsilon$  (cf. the proof of Satz 6.3 in [1]), we obtain

$$\begin{aligned} \beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) &\leq \beta(\alpha, (\mathcal{P}'_0)_{\tilde{\mu}}, (\mathcal{P}'_1)_{\tilde{\mu}}) = \\ &\inf \left\{ \alpha k + \frac{1-k}{2} + \frac{1}{2} d_k(\text{con } (\mathcal{P}'_1)_{\tilde{\mu}}, \text{con } (\mathcal{P}'_0)_{\tilde{\mu}}) \right\} = \\ &\inf \left\{ \alpha k + \frac{1-k}{2} + \frac{1}{2} d_k((\text{con } \mathcal{P}'_1)_{\tilde{\mu}}, (\text{con } \mathcal{P}'_0)_{\tilde{\mu}}) \right\} \leq \\ &\inf \left\{ \alpha k + \frac{1-k}{2} + \frac{1}{2} d_k(\text{con } \mathcal{P}'_1, \text{con } \mathcal{P}'_0) \right\} + \epsilon = \\ &\inf \left\{ \alpha k + \frac{1-k}{2} + \frac{1}{2} d_k(\text{con } \mathcal{P}_1, \text{con } \mathcal{P}_0) \right\} + \epsilon = \\ &\beta(\alpha, \mathcal{P}_0, \mathcal{P}_1) + \epsilon, \quad \forall \epsilon > 0. \end{aligned}$$

Observe that by Proposition 9 the assumptions of Theorem 10 are fulfilled for  $\mathcal{P}'_i = E(\mathcal{P}_i)$ . Without assuming that  $\mathcal{P}_0 \cup \mathcal{P}_1$  is dominated we obtain:

**Theorem 11.** *If for each  $P \in \mathcal{P}_i$  there exists a  $Q_P \in E(P) \cap \overline{\text{con } \mathcal{P}'_i}$ ,  $i = 0, 1$ , such that  $\mathcal{A}_0$  is sufficient for  $M = \{Q_P \mid P \in \mathcal{P}_0 \cup \mathcal{P}_1\}$ , then  $\beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) = \beta(\alpha, \mathcal{P}_0, \mathcal{P}_1)$ .*

**Proof.** Let  $\varphi \in \Phi_\alpha(\mathcal{P}'_0, \mathcal{A})$  and define  $\psi = E_\mu(\varphi \mid \mathcal{A}_0)$ ,  $\mu \in M$ . Then for  $P \in \mathcal{P}_0$   $E_P \psi = E_{Q_P} \psi = E_{Q_P} \varphi \leq \alpha$  which implies  $\psi \in \Phi_\alpha(\mathcal{P}'_0, \mathcal{A})$ . Furthermore, for  $P \in \mathcal{P}_1$ , and  $Q \in \mathcal{P}'_1 \cap E(P)$  it holds

that  $E_Q \psi = E_P \psi = E_{Q_P} \psi = E_{Q_P} \varphi$  and, therefore,  $\inf_{Q \in \mathcal{P}'_1} E_Q \varphi \leq \inf_{Q \in \mathcal{P}'_1} E_Q \psi$ . This implies that  $\beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1) = \beta(\alpha, \mathcal{P}_0, \mathcal{P}_1)$ .

**Remark 2.**

(a) The condition of Theorem 11 corresponds to an assumption made by Hajek [5] in the case of estimation in the presence of nuisance parameters. The conclusion of Theorem 11 could be strengthened to

$$\beta(\alpha, \mathcal{P}'_0, E(P) \cap \overline{\text{con}} P'_1) = \beta(\alpha, \mathcal{P}_0, P) \quad \text{for all } P \in \mathcal{P}_1.$$

(b) If  $\mathcal{P}_i = \{P_i\}$ ,  $i = 0, 1$ , then the condition that there exist  $Q_P \in E(P_i) \cap \overline{\text{con}} \mathcal{P}'_i$ ,  $i = 0, 1$ , such that  $\mathcal{A}_0$  is sufficient for  $\{Q_{P_0}, Q_{P_1}\}$  is equivalent to the assumption that there exists a  $\mu \in M(\mathcal{X}, \mathcal{A})$  with  $P_0, P_1 \ll \mu | \mathcal{A}_0$  and  $Q_i = P_{i, \mu} \in \overline{\text{con}} \mathcal{P}'_i$ ,  $i = 0, 1$  (cf. also Corollary 2). The determination of a maximin-test is simplified in the presence of least favourable pairs. In the literature there are three different definitions of least favourable pairs for the testproblem  $\mathcal{P}'_0, \mathcal{P}'_1$ .

Let  $P_i \in \overline{\text{con}} \mathcal{P}'_i$ ,  $i = 0, 1$ , then

(a)  $(P_0, P_1) \in \text{LF}_\alpha(\mathcal{P}'_0, \mathcal{P}'_1)$  iff  $\beta(\alpha, P_0, P_1) = \beta(\alpha, \mathcal{P}'_0, \mathcal{P}'_1)$ ,  
 $(P_0, P_1) \in \text{LF}(\mathcal{P}'_0, \mathcal{P}'_1)$  iff  $(P_0, P_1) \in \bigcap_{\alpha \in [0, 1]} \text{LF}_\alpha(\mathcal{P}'_0, \mathcal{P}'_1)$  (cf. Baumann [1]).

(b)  $(P_0, P_1) \in \widetilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1)$  iff there exists  $\pi \in \frac{dP_1}{dP_0}$  with  $P^\pi \leq_{\text{st}} P_0^\pi$  for  $P \in \mathcal{P}'_0$ ,  $Q^\pi \geq_{\text{st}} P_1^\pi$  for  $Q \in \mathcal{P}'_1$  where  $\leq_{\text{st}}$  is the stochastic order and  $\frac{dP_1}{dP_0} = \left\{ \frac{f_1}{f_0}, f_i = \frac{dP_i}{dP_\mu}, i = 0, 1, \mu \text{ dominates } P_i, i = 0, 1 \right\}$  (cf. Huber, Strassen [6], Rieder [11]).

(c)  $(P_0, P_1) \in \text{LF}'_\alpha(\mathcal{P}'_0, \mathcal{P}'_1)$  iff  $\beta_\alpha(P_0, P_1) \leq \beta_\alpha(Q_0, Q_1)$  for all  $Q_i \in \overline{\text{con}} \mathcal{P}'_i$ ,  $i = 0, 1$ ,  $(P_0, P_1) \in \text{LF}'(\mathcal{P}'_0, \mathcal{P}'_1)$  iff  $(P_0, P_1) \in \bigcap_{\alpha \in [0, 1]} \text{LF}'_\alpha(\mathcal{P}'_0, \mathcal{P}'_1)$  (cf. Lehmann [7], pg. 325).

The following remark is concerned with connections of these notions of least favourable pairs and with methods to find least favourable pairs.

**Remark 3.**

(1a)  $(P_0, P_1) \in \text{LF}_\alpha(\mathcal{P}'_0, \mathcal{P}'_1)$  iff there exists a most powerful level  $\alpha$  test for  $P_0, P_1$  which is maximin-test for  $\mathcal{P}'_0, \mathcal{P}'_1$  at level  $\alpha$ .

(1b)  $(P_0, P_1) \in \tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1)$  iff there exists a  $\pi \in \frac{dP_1}{dP_0}$  such that  $\varphi_{\pi, \alpha}$  (the LQ-test at level  $\alpha$  which is constant on the randomized region) is maximin-test for  $\mathcal{P}'_0, \mathcal{P}'_1$  for each  $\alpha \in [0, 1]$ .

(1c) From (1b) follows  $\tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1) \subset \text{LF}(\mathcal{P}'_0, \mathcal{P}'_1)$  (there is no equality in general). Equality holds if for instance the distribution of  $\pi \in \frac{dP_1}{dP_0}$  is nonatomic under  $\mathcal{P}'_i, i = 0, 1$ , for all  $(P_0, P_1) \in \text{LF}(\mathcal{P}'_0, \mathcal{P}'_1)$ .

(1d) If  $\tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1) \neq \emptyset$ , then  $\text{LF}(\mathcal{P}'_0, \mathcal{P}'_1) = \text{LF}'(\mathcal{P}'_0, \mathcal{P}'_1)$  (cf. Rieder [11], Proposition 2.2).

(2)  $\tilde{\text{LF}}((\mathcal{P}'_0)^{(n)}, (\mathcal{P}'_1)^{(n)}) = (\tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1))^{(n)}$ , where  $(\mathcal{P}'_i)^{(n)} = \{P^{(n)}: P \in \mathcal{P}'_i\}$  (cf. Huber, Strassen [6], Corollary 4).

(3a)  $(P_0, P_1) \in \tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1) \Rightarrow d_k(P_0, P_1) = d_k(\text{con } \mathcal{P}'_0, \text{con } \mathcal{P}'_1)$  for all  $k \geq 0$ .

(3b) If  $\tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1) \neq \emptyset$ ,  $P_i \in \overline{\text{con } \mathcal{P}'_i}, i = 0, 1$ , with  $d_k(P_0, P_1) = d_k(\text{con } \mathcal{P}'_0, \text{con } \mathcal{P}'_1)$  for all  $k \geq 0$ , then  $(P_0, P_1) \in \tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1)$ . For similar facts concerning  $\text{LF}(\mathcal{P}'_0, \mathcal{P}'_1)$ , cf. Reinhardt [10].

(4) Elements of  $\tilde{\text{LF}}(\mathcal{P}'_0, \mathcal{P}'_1)$  can also be determined by minimization of certain different distance measures containing for example the measure of divergence of Csiszár [2]. Let  $\varphi: [0, 1] \rightarrow \mathbf{R}^1$  be twice continuously differentiable with  $\varphi'' > 0$  and define:

$$H(P, Q) = \int \varphi\left(\frac{dP}{d(P+Q)}\right) d(P+Q)$$

for probability measures  $P, Q$  on  $(\mathcal{X}, \mathcal{A})$ . Then by a slight modification of Theorem 6.1 of Huber, Strassen [6]:

$$H(\mathscr{P}, \mathscr{Q}) = \inf \{H(P, Q), P \in \mathscr{P}, Q \in \mathscr{Q}\}.$$

$$(a) (P_0, P_1) \in \widetilde{\text{LF}}(\mathscr{P}'_0, \mathscr{P}'_1) \Rightarrow H(P_0, P_1) = H(\text{con } \mathscr{P}'_0, \text{con } \mathscr{P}'_1).$$

(b) If  $P_i \in \overline{\text{con } \mathscr{P}'_i}$ ,  $i = 0, 1$ ,  $\widetilde{\text{LF}}(\mathscr{P}'_0, \mathscr{P}'_1) \neq \phi$  and  $H(P_0, P_1) = H(\text{con } \mathscr{P}'_0, \text{con } \mathscr{P}'_1)$ , then  $(P_0, P_1) \in \widetilde{\text{LF}}(\mathscr{P}'_0, \mathscr{P}'_1)$ .

Returning to our testproblem  $\mathscr{P}'_0 \subset E(\mathscr{P}_0)$ ,  $\mathscr{P}'_1 \subset E(\mathscr{P}_1)$  we have

**Theorem 12.** *If  $(P_0, P_1) \in \text{LF}_\alpha(\mathscr{P}_0, \mathscr{P}_1)$  ( $\widetilde{\text{LF}}(\mathscr{P}_0, \mathscr{P}_1)$ ) and if there exist  $Q_i \in \overline{\text{con } \mathscr{P}'_i} \cap E(P_i)$ ,  $i = 0, 1$ , such that  $\mathscr{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ , then*

$$(a) (Q_0, Q_1) \in \text{LF}_\alpha(\mathscr{P}'_0, \mathscr{P}'_1) (\widetilde{\text{LF}}(\mathscr{P}'_0, \mathscr{P}'_1)).$$

(b) *There is a most powerful level  $\alpha$  test  $\varphi_\alpha$  for  $P_0, P_1$  which is maximin-test at level  $\alpha$  for  $\mathscr{P}'_0, \mathscr{P}'_1$ . ( $\varphi_{\pi, \alpha}$  is maximin-test at level  $\alpha$  for  $\mathscr{P}'_0, \mathscr{P}'_1$  where  $\pi \in \frac{dP_1}{dP_0}$ ).*

**Proof.**

I. Let  $(P_0, P_1) \in \text{LF}_\alpha(\mathscr{P}_0, \mathscr{P}_1)$ .

(1) By Remark 3 (1a) there is a most powerful level  $\alpha$  test  $\varphi_\alpha$  for  $P_0, P_1$  which is maximin-test for  $\mathscr{P}_0, \mathscr{P}_1$ . Therefore,  $\varphi_\alpha \in \Phi_\alpha(\mathscr{P}'_0, \mathscr{A})$ .

(2) Let  $\varphi \in \Phi_\alpha(\mathscr{P}'_0, \mathscr{A})$  and define  $\psi = E_{\{Q_0, Q_1\}}(\varphi | \mathscr{A}_0)$ . Then  $\inf_{Q \in \mathscr{P}'_1} E_Q \varphi \leq E_{Q_1} \varphi$ , since  $Q_1 \in \overline{\text{con } \mathscr{P}'_1}$ , and  $E_{Q_1}(\varphi) = E_{Q_1} \psi = E_{P_1} \psi$ .

Clearly  $\psi \in \Phi_\alpha(\mathscr{P}_0, \mathscr{A}_0)$  and, therefore,

$$E_{P_1} \psi \leq E_{P_1} \varphi_\alpha = \inf_{P \in \mathscr{P}_1} E_P \varphi_\alpha = \inf_{Q \in \mathscr{P}'_1} E_Q \varphi_\alpha.$$

This implies  $\beta(\alpha, Q_0, Q_1) = \beta(\alpha, P_0, P_1) = \beta(\alpha, \mathscr{P}'_0, \mathscr{P}'_1)$ , i.e.  $(Q_0, Q_1) \in \text{LF}_\alpha(\mathscr{P}'_0, \mathscr{P}'_1)$  and clearly  $\varphi_\alpha$  is a maximin-test for  $\mathscr{P}'_0, \mathscr{P}'_1$ .

II. The case  $(P_0, P_1) \in \widetilde{\text{LF}}(\mathscr{P}'_0, \mathscr{P}'_1)$  is similar.

**Corollary 3.** *If  $(P_0, P_1) \in \text{LF}(\mathscr{P}_0, \mathscr{P}_1)$  and if there exist  $Q_i \in \overline{\text{con } \mathscr{P}'_i} \cap E(P_i)$ ,  $i = 0, 1$ , such that  $\mathscr{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ , then*

$$(Q_0^{(n)}, Q_1^{(n)}) \in \widetilde{\text{LF}}((\mathcal{P}'_0)^{(n)}, (\mathcal{P}'_1)^{(n)}).$$

**Proof.** Observe that

$$\overline{\text{con}}(E(P) \times E(Q)) \subset E(P \times Q)$$

for  $P, Q \in M_1(\mathcal{X}, \mathcal{A})$  and, therefore,  $Q_i^{(n)} \in E(P_i^{(n)})$ ,  $i = 0, 1$ . By Remark 3.2  $(P_0^{(n)}, P_1^{(n)}) \in \widetilde{\text{LF}}(\mathcal{P}_0^{(n)}, \mathcal{P}_1^{(n)})$  and, furthermore,  $\mathcal{A}_0^{(n)}$  (the  $n$ -fold product of  $\mathcal{A}_0$ ) is sufficient for  $\{Q_0^{(n)}, Q_1^{(n)}\}$ . Therefore, Corollary 3 follows from Theorem 12.

**Corollary 4.** Let  $(P_0, P_1) \in \text{LF}_\alpha(\mathcal{P}_0, \mathcal{P}_1)$  and let  $\varphi_\alpha$  be a most powerful level  $\alpha$  test for  $P_0, P_1$ , which is maximin for  $\mathcal{P}_0, \mathcal{P}_1$ . Then  $\varphi_\alpha$  is maximin-test at level  $\alpha$  for  $E(\mathcal{P}_0), E(\mathcal{P}_1)$ .

#### 4. UNIFORMLY MOST POWERFUL TESTS

Again let  $\mathcal{P}_i \subset M_1(\mathcal{X}, \mathcal{A}_0)$  and  $\mathcal{P}'_i \subset E(\mathcal{P}_i)$ ,  $i = 0, 1$ , and let  $E(P) \neq \emptyset$ ,  $\forall P \in \mathcal{P}_0 \cup \mathcal{P}_1$ . For  $P, Q \in M_1(\mathcal{X}, \mathcal{A}_0)$  let  $\Phi_\alpha^*(P, Q)$  denote the set of most powerful level  $\alpha$  tests for  $P, Q$ .

**Theorem 13.**

- (a) Let  $\varphi_0$  be a UMP-test for  $\mathcal{P}_0, \mathcal{P}_1$  at level  $\alpha$ ,
- (b) Let there exist  $P_0 \in \mathcal{P}_0$  such that for all  $P_1 \in \mathcal{P}_1$ ,  $\varphi_0 \in \Phi_\alpha^*(P_0, P_1)$ ,
- (c) For all  $Q_1 \in \mathcal{P}'_1$  let there exist a  $Q_0 \in \overline{\text{con}} \mathcal{P}'_0 \cap E(P_0)$  such that  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ .

Then  $\varphi_0$  is a UMP-test for  $\mathcal{P}'_0, \mathcal{P}'_1$  at level  $\alpha$ .

**Proof.** Clearly  $\varphi_0 \in \Phi_\alpha(\mathcal{P}'_0, \mathcal{A})$ . Let  $\varphi \in \Phi_\alpha(\mathcal{P}'_0, \mathcal{A})$  and let  $Q_1 \in \mathcal{P}'_1$ . Then there exists  $P_1 \in \mathcal{P}_1$  such that  $Q_1 \in E(P_1)$  and so by (c) there is a  $Q_0 \in \overline{\text{con}} \mathcal{P}'_0 \cap E(P_0)$  such that  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ . Define  $\psi = E_{\{Q_0, Q_1\}}(\varphi | \mathcal{A}_0)$ , then  $E_{P_0} \psi = E_{Q_0} \psi = E_{Q_0} \varphi \leq \alpha$  and, therefore, (b) implies

$$E_{Q_1} \varphi = E_{Q_1} \psi = E_{P_1} \psi \leq E_{P_1} \varphi_0 = E_{Q_1} \varphi_0.$$

This yields that  $\varphi_0$  is UMP at level  $\alpha$  for  $\mathcal{P}'_0, \mathcal{P}'_1$ .

**Corollary 5.** Let  $P_i \in M_1(\mathcal{X}, \mathcal{A}_0)$  with  $E(P_i) \neq \emptyset$ ,  $i = 0, 1$ , and  $\varphi_0 \in \Phi_\alpha^*(P_0, P_1)$ . Then

- (a)  $\varphi_0$  is UMP-test at level  $\alpha$  for  $E(P_0), E(P_1)$ ,  
 (b) If  $\varphi^*$  is a UMP-test at level  $\alpha$  for  $E(P_0), E(P_1)$  and if  $E_{Q_1} \varphi_0 < 1$  for all  $Q_1 \in E(P_1)$ , then  $\varphi^* = E_Q(\varphi^* | \mathcal{A}_0)[Q]$  for all  $Q = \frac{1}{2}(Q_0 + Q_1)$ ,  $Q_i \in E(P_i)$ ,  $i = 0, 1$ .

**Proof.**

(a) Let  $P_0 = aP'_0 + (1-a)P''_0$  be a decomposition as in Proposition 8 and let  $Q_1 \in E(P_1)$ . Then  $Q_0 := ahQ_1 + (1-a)Q_0$  (with  $\tilde{Q}_0 \in E(P''_0)$  and  $h$  a version of  $\frac{dP'_0}{dP_1}$ ) is an element of  $E(P_0)$  such that  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ . So (a) is implied by Theorem 13.

(b) For all  $Q_0 \in E(P_0)$  we can find a  $Q_1 \in E(P_1)$  (as in (a)) such that  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ . Therefore,  $\varphi_0 \in \Phi_\alpha^*(Q_0, Q_1)$  and, therefore, also  $\varphi^* \in \Phi_\alpha^*(Q_0, Q_1)$ . Since  $\beta = E_{Q_1} \varphi_0 < 1$  we have  $E_{Q_0} \varphi^* = \alpha$ . This implies that  $\varphi^*$  is a UMP-test at level  $\alpha$  for  $E(P_0)$  against  $\frac{1}{2}E(P_0) + \frac{1}{2}E(P_1) = E\left(\frac{1}{2}(P_0 + P_1)\right)$  by Proposition 6, and, therefore,  $E_Q \varphi^* = \frac{\alpha + \beta}{2}$  for all  $Q \in E\left(\frac{P_0 + P_1}{2}\right)$ . Proposition 5 implies that  $\varphi^* = E_Q(\varphi^* | \mathcal{A}_0)[Q]$  for all  $Q \in E\left(\frac{1}{2}(P_0 + P_1)\right)$ .

**Remark 4.** Corollary 5 generalizes a result of Fraser [4], Theorem 2, which is concerned with the case of nuisance parameters.

## 5. EXAMPLES

(1) Let  $a \leq b$  and  $0 \leq \alpha \leq \beta$ ,  $\alpha + \beta \leq 1$  and consider the test-problem

$$\mathcal{P}'_0 = \{P^{(n)} \mid P \in M_1(\mathbb{R}^1, \mathcal{B}^1), P(-\infty, a] \leq \alpha, P(a, b] \leq \alpha\},$$

$$\mathcal{P}'_1 = \{P^{(n)} \mid P \in M_1(\mathbf{R}^1, \mathcal{B}^1), P(-\infty, a] \geq \alpha, P(a, b] \geq \beta\}.$$

Let  $T: \mathbf{R}^n \rightarrow \mathbf{R}^2$ .  $T(x) = (s_1(x), s_2(x))$ , where  $s_1(x) = \sum_{i=1}^n 1_{(-\infty, a]}(x_i)$ ,  $s_2(x) = \sum_{i=1}^n 1_{(a, b]}(x_i)$ . Let  $\mathcal{A}_0 = \mathcal{A}(T)$  the  $\sigma$ -algebra induced by  $T$  and let  $\mathcal{P}_i$  denote the restriction of  $\mathcal{P}'_i$  on  $\mathcal{A}_0$ ,  $i = 0, 1$ , so that we have the situation considered in Sections 3, 4.

To determine  $\mathcal{P}_i$  let  $A_0 = (-\infty, a]$ ,  $A_1 = (a, b]$ ,  $A_2 = (b, \infty)$  and use the representation

$$\mathcal{P}'_0 = \left\{ \left( \sum_{i=0}^2 \alpha_i P_i \right)^{(n)}, P_i \in M_1(A_i, \mathcal{A}_i \mathcal{P}_1), i = 0, 1, 2, \right.$$

$$\left. 0 \leq \alpha_i \leq 1, \sum \alpha_i = 1 \text{ and } \alpha_0 \leq \alpha, \alpha_1 \leq \alpha \right\},$$

$$\mathcal{P}'_1 = \left\{ \left( \sum \alpha_i P_i \right)^{(n)}, P_i \in M_1(A_i, \mathcal{A}_i \mathcal{P}_1), i = 0, 1, 2, \right.$$

$$\left. 0 \leq \alpha_i \leq 1, \sum \alpha_i = 1, \alpha_0 \geq \alpha, \alpha_1 \geq \beta \right\}.$$

If  $Q = \left( \sum \alpha_i P_i \right)^{(n)} \in \mathcal{P}'_0 \cup \mathcal{P}'_1$ , then

$$Q(s_1 = k, s_2 = m) = \binom{n}{k, m} \alpha_0^k \alpha_1^m \alpha_2^{n-(k+m)};$$

so  $Q/\mathcal{A}_0 = Q(\alpha_0, \alpha_1)$  and

$$\mathcal{P}_0 = \{Q(\alpha_0, \beta_0) \mid \alpha_0 \leq \alpha, \beta_0 \leq \alpha\},$$

$$\mathcal{P}_1 = \{Q(\alpha_1, \beta_1) \mid \alpha_1 \geq \alpha, \beta_1 \geq \beta\}.$$

Let  $Q_i = Q_i(\alpha_i, \beta_i) \in \mathcal{P}_i$ ,  $i = 0, 1$ , then

$$\frac{Q_1(s_1 = k, s_2 = m)}{Q_0(s_1 = k, s_2 = m)} =$$

$$\left( \frac{\alpha_1(1 - (\alpha_0 + \beta_0))}{\alpha_0(1 - (\alpha_1 + \beta_1))} \right)^k \left( \frac{\beta_1(1 - (\alpha_0 + \alpha_0))}{\beta_0(1 - (\alpha_1 + \beta_1))} \right)^m.$$

From this we easily obtain that

$$(Q(\alpha, \alpha), Q(\alpha, \beta)) \in \tilde{L}F(\mathcal{P}_0, \mathcal{P}_1)$$

with most powerful level  $\alpha$  test of the type

$$\varphi_0(x) = \begin{cases} 1 & \text{if } \left(\frac{\beta}{\alpha}\right)^{s_1(x)} \left(\frac{1-2\alpha}{1-(\alpha+\beta)}\right)^{s_1(x)+s_2(x)} > k_\alpha, \\ 0 & \leq k_\alpha. \end{cases}$$

Clearly

$$Q_0 = (\alpha P_0 + \alpha P_1 + (1-2\alpha)P_2)^{(n)},$$

$$Q_1 = (\alpha P_0 + \beta P_1 + (1-(\alpha+\beta))P_2)^{(n)}$$

define extensions of  $Q(\alpha, \alpha), Q(\alpha, \beta)$  in  $\mathcal{P}'_i$ , such that  $\mathcal{A}_0$  is sufficient for  $\{Q_0, Q_1\}$ . ( $P_i$  are any elements of  $M_1(A_i, A_i, \mathcal{L}_1)$ ).

So by Theorem 12  $\varphi_0$  is a maximin-test at level  $\alpha$  for  $\mathcal{P}'_0, \mathcal{P}'_1$ . Clearly no UMP-test exists in this situation.

(2) Let  $P_{(\mu, \sigma^2)} = \bigotimes_{i=1}^n N(\mu, \sigma^2)$ ,  $\mu \in \mathbf{R}^1$ ,  $\sigma^2 > 0$  and consider the testproblem:

(a)  $\mathcal{P}'_0 = \{P_{\mu, \sigma^2} \mid \sigma^2 \leq \sigma_0^2\}$ ,  $\mathcal{P}'_1 = \{P_{\mu, \sigma^2} \mid \sigma^2 \geq \sigma_1^2\}$ , where  $\sigma_0^2 < \sigma_1^2$ . If  $s^2(x) = \sum_{i=1}^n (x_i - \bar{x})^2$ ,  $\mathcal{A}_0 = \mathcal{A}(s^2)$ , then  $P_{\mu, \sigma^2}^{s^2} = h_\sigma \lambda^n$  has monotone likelihood ratio in  $\sigma^2$  so that

$$\varphi_0 = \begin{cases} 1, & s^2(x) > k_\alpha, \\ 0, & s^2(x) \leq k_\alpha \end{cases}$$

is a UMP-test at level  $\alpha$  for

$$\mathcal{P}_0 = \{h_\sigma \lambda^n \mid \sigma^2 \leq \sigma_0^2\}, \mathcal{P}_1 = \{h_\sigma \lambda^n \mid \sigma^2 \geq \sigma_1^2\}. P_0 = h_{\sigma_0^2} \lambda^n$$

satisfies condition (b) of Theorem 13. For condition (c) let  $\mu \in \mathbf{R}^1$ ,  $\sigma^2 \geq \sigma_1^2$  and  $Q_1 = P_{(\mu, \sigma^2)}$ . We are looking for

$$Q_0 \in \overline{\text{con}} \{P_{\mu', \sigma_0^2} \mid \mu' \in \mathbf{R}^1\},$$

such that  $s^2$  is sufficient for  $\{Q_0, Q_1\}$ . Let  $P_{\mu', \sigma_0^2}$  be the density of  $P_{\mu', \sigma_0^2}$  with respect to  $\lambda^n$ , then



$$P_{\mu', \sigma_0^2}(x) = A(\sigma_0^2) \exp\left(-\frac{s^2}{2\sigma_0^2}\right) \exp\left(-\frac{n(\bar{x} - \mu')^2}{2\sigma_0^2}\right).$$

Using

$$N(\bar{x}, \frac{\sigma_0^2}{n}) * N(\mu - \bar{x}, \frac{\sigma^2}{n} - \frac{\sigma_0^2}{n}) = N(\mu, \frac{\sigma^2}{n})$$

(\* denoting convolution), we obtain with  $\lambda_0 = N(\mu - \bar{x}, \frac{\sigma^2}{n} - \frac{\sigma_0^2}{n})$ ,

$Q_0 := \int P_{\mu', \sigma_0^2} d\lambda_0(\mu')$  has  $\lambda^n$ -density

$$A(\sigma_0^2, \sigma^2) \exp\left(-\frac{s^2}{2\sigma_0^2}\right) e^{-\frac{n(\bar{x} - \mu)^2}{2\sigma^2}},$$

which shows that  $s^2$  is sufficient for  $\{Q_0, Q_1\}$ .

Theorem 13 implies that  $\varphi_0$  is a UMP-test at level  $\alpha$  for  $\mathcal{P}'_0, \mathcal{P}'_1$ . So in this well-known case we obtain an explanation why to choose the mixing measure  $\lambda_0$ .

(b) Similarly for the testing problem  $\mathcal{P}'_0 = \{P_{\mu, \sigma^2} \mid \sigma_0^2 \leq \sigma^2 \leq K\}$  against  $\mathcal{P}'_1 = \{P_{\mu, \sigma^2} \mid \sigma^2 < \sigma_0^2\}$  we obtain that  $s^2$  is sufficient for  $\{Q_{\sigma^2} \mid 0 < \sigma^2 \leq K\}$ , where  $Q_{\sigma^2} = \int P_{\mu, \sigma^2} d\lambda_{\sigma^2}(\mu)$  and where  $\lambda_{\sigma^2} = N(0, \frac{K - \sigma^2}{n})$ . So by Theorem 11

$$\varphi_0 = \begin{cases} 1, & s^2(x) > k_\alpha, \\ 0, & s^2(x) \leq k_\alpha \end{cases}$$

yields a maximin-test for  $\mathcal{P}'_0, \mathcal{P}'_1$  (which is independent of  $K$ ).

(3) Let  $P_{\alpha, \beta} = f_{\alpha, \beta} \mu$  with  $(\alpha, \beta) \in \Theta$  and let  $A$  be the projection of  $\Theta$  onto the first component. Assume

$$(a) f_{\alpha, \beta}(x) = f_\alpha(T(x))g_{\alpha, \beta}(x), \quad \forall x \in \mathcal{X}, \quad \alpha \in A, \quad \beta \in \Theta_\alpha.$$

(b) For each  $\alpha \in A$  there is a probability measure  $\lambda_\alpha$  on  $\Theta_\alpha$ , such that  $\int_B g_{\alpha, \beta}(x) d\lambda_\alpha(\beta) = Q(B)$ ,  $B \in \mathcal{A}$ , is independent of  $\alpha \in A$ , then by Theorem 11 testproblems with respect to  $\alpha$  can be reduced to the test-

problems for the distributions of  $T$ , when considering the maximin-risk. Examples are:  $g_{\alpha,\beta}(x) = g_\beta(x)$  with  $\lambda_\alpha = \epsilon_{\{\beta_0\}}$ ,  $g_{\alpha,\beta}(x) = h_{\alpha-\beta}(x)$  with  $\lambda_\alpha = \epsilon_{\{-\alpha\}}$  and  $g_{\alpha,\beta}(x) = h_{\frac{\alpha}{\beta}}(x)$  with  $\lambda_\alpha = \epsilon_{\{\alpha\}}$ .

(4) Let  $G$  denote a finite group of order  $\gamma$  consisting of  $(\mathcal{A}, \mathcal{A})$ -measurable transformations  $g: \mathcal{X} \rightarrow \mathcal{X}$  and introduce  $\mathcal{A}_0$  as the sub- $\sigma$ -algebra  $\mathcal{A}_0$  of  $\mathcal{A}$  consisting of all  $G$ -invariant sets belonging to  $\mathcal{A}$ . If  $P_i$  are probability measures on  $\mathcal{A}$ ,  $\mathcal{P}_i$ ,  $i = 0, 1$ , is defined to be the family  $\{P_i^g | g \in G\}$ ,  $i = 0, 1$ . A UMP-test for  $P_0 | \mathcal{A}_0, P_1 | \mathcal{A}_0$  at level  $\alpha$  is in this case according to Theorem 11 a maximin-test for  $\mathcal{P}_0, \mathcal{P}_1$  at level  $\alpha$ , since  $Q_i = \frac{1}{\gamma} \sum_{g \in G} P_i^g \in \text{con } \mathcal{P}_i$ ,  $i = 0, 1$ , and  $E_{Q_i}(I_A | \mathcal{A}_0) = \frac{1}{\gamma} \sum_{g \in G} I_A \circ g$ ,  $A \in \mathcal{A}$ ,  $i = 0, 1$ . Especially the version of the sign in Lehmann's book, p. 219–220, is a maximin-test at level  $\alpha$ .

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## 1. INTRODUCTION AND SUMMARY

In some recent papers (Rasch [5], Rasch and Schinke [6] [7]) we investigated the problem of whether slightly modified statistics in non-linear regression models which are known to be asymptotically normally distributed can be used to test hypotheses on model parameters also for finite  $n$ . Further, we wanted to know what the smallest  $n$  is for the first kind risk not to deviate by more than 10% from the nominal one. We especially considered the exponential regression, i.e. the model with i. i. d. errors  $\varepsilon_i$  with zero expectation:

$$(1) \quad y_i = \alpha + \beta e^{\gamma x_i} + \varepsilon_i, \quad i = 1, \dots, n > 3, \beta < 0, \gamma < 0$$

(random variables are underlined). We found that the tests for confidence estimations work very well even if  $n$  is as small as 4 – not only if the  $\varepsilon_i$  are normally distributed but also if we use non-normal error distributions with special values of skewness  $\gamma_1$  and kurtosis  $\gamma_2$ . But the limitation of